

Let  $Y = (Y_1, \dots, Y_n)$  be a random vector with mean  $\mu_Y$  and covariance matrix  $\Sigma_Y$ . Let  $X = (X_1, \dots, X_M)$  be a random vector with mean  $\mu_X$  and covariance matrix  $\Sigma_X$ . Let  $K = (K_1, \dots, K_M)$  be a random vector with mean  $\mu_K$  and covariance matrix  $\Sigma_K$ . Let  $M = (M_1, \dots, M_M)$  be a random vector with mean  $\mu_M$  and covariance matrix  $\Sigma_M$ . Let  $H = (H_1, \dots, H_M)$  be a random vector with mean  $\mu_H$  and covariance matrix  $\Sigma_H$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

3. FUNCTIONAL ADDITIVE MODELING

Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively. Let  $b_k$  and  $b_{km}$  be constants. Let  $E(Y - \mu_Y | k) = b_k$  and  $E(m | k) = b_{km}$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

$$E(Y - \mu_Y | k) = b_k \quad E(m | k) = b_{km} \quad (1)$$

Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively. Let  $b_k$  and  $b_{km}$  be constants. Let  $E(Y - \mu_Y | k) = b_k$  and  $E(m | k) = b_{km}$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

$$f_k(\cdot) \quad f_{km}(\cdot) \quad k, m = 1, \dots, M \quad (2)$$

$$E(Y|X) = \mu_Y + \sum_{k=1}^{\infty} f_k(k) \quad (3)$$

$$E(Y(t)|X) = \mu_Y(t) + \sum_{k=1}^{\infty} \sum_{m=1}^{\infty} f_{km}(k) m(t), \quad (4)$$

$$E f_k(k) = \dots, \quad k = 1, \dots, M \quad (5)$$

$$E f_{km}(k) = \dots, \quad k = 1, \dots, M, m = 1, \dots, M \quad (6)$$

Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively. Let  $b_k$  and  $b_{km}$  be constants. Let  $E(Y - \mu_Y | k) = b_k$  and  $E(m | k) = b_{km}$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

$$E(Y - \mu_Y | k) = E\{E(Y - \mu_Y | X) | k\}$$

$$= E\left\{\sum_{j=1}^{\infty} f_j(j) | k\right\} = f_k(k), \quad (7)$$

$$E(m | k) = E\{E(m | X) | k\}$$

$$= E\left\{\sum_{j=1}^{\infty} f_{jm}(j) | k\right\} = f_{km}(k). \quad (8)$$

$$E(m | k) \quad f_k(k) = E(Y - \mu_Y | k) \quad f_{km}(k) =$$

Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively. Let  $b_k$  and  $b_{km}$  be constants. Let  $E(Y - \mu_Y | k) = b_k$  and  $E(m | k) = b_{km}$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively. Let  $b_k$  and  $b_{km}$  be constants. Let  $E(Y - \mu_Y | k) = b_k$  and  $E(m | k) = b_{km}$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

4. FITTING OF FUNCTIONAL ADDITIVE MODELS

Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively. Let  $b_k$  and  $b_{km}$  be constants. Let  $E(Y - \mu_Y | k) = b_k$  and  $E(m | k) = b_{km}$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

principal analysis by conditional expectation (E)

$$E \quad K \quad (9)$$

Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively. Let  $b_k$  and  $b_{km}$  be constants. Let  $E(Y - \mu_Y | k) = b_k$  and  $E(m | k) = b_{km}$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively. Let  $b_k$  and  $b_{km}$  be constants. Let  $E(Y - \mu_Y | k) = b_k$  and  $E(m | k) = b_{km}$ . Let  $f_k(\cdot)$  and  $f_{km}(\cdot)$  be functions of  $k$  and  $m$ , respectively.

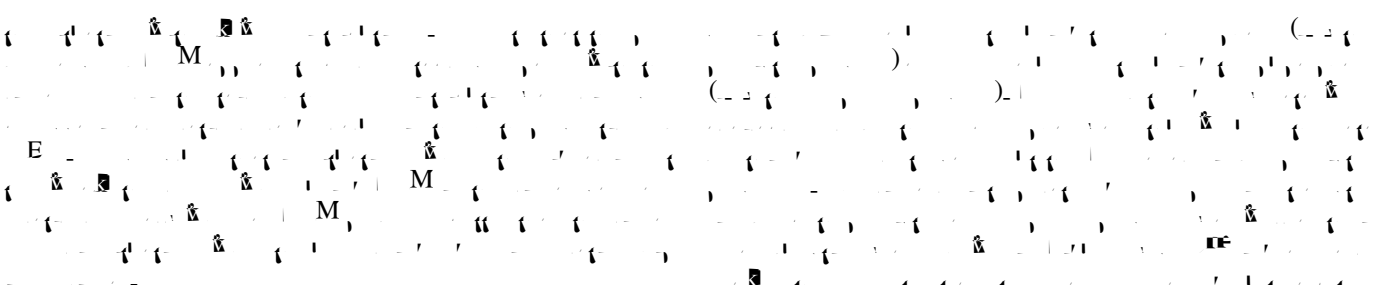
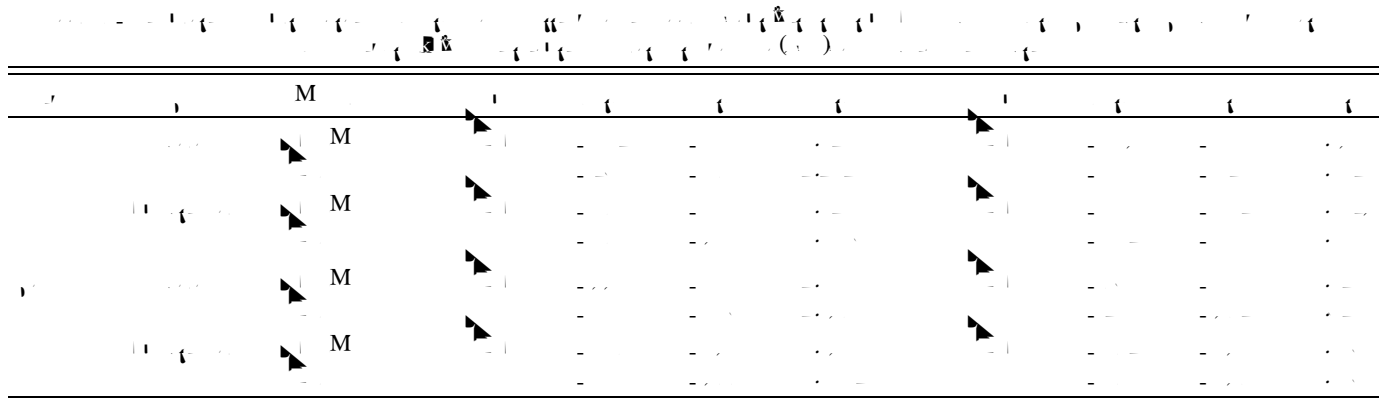
$$\sum_{i=1}^n K \left( \frac{\hat{ik} - X}{h_k} \right) \{Y_i - (X - \hat{ik})\} \quad (10)$$

$$\hat{f}_k(x) = \hat{(x)} - \bar{Y} \quad h_k$$

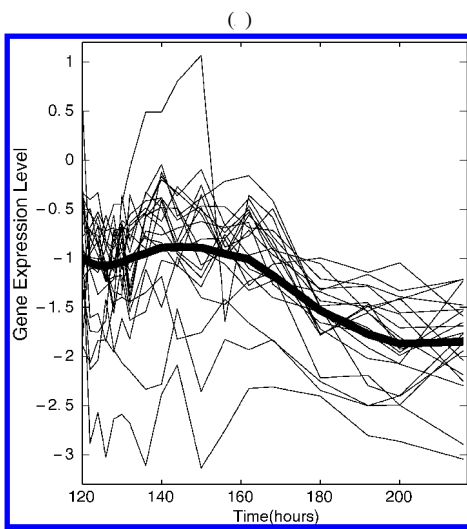
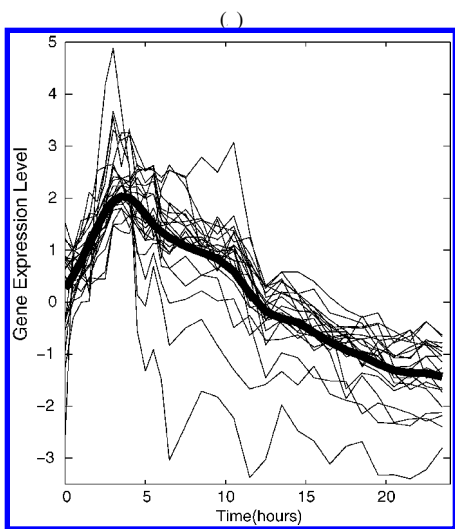
$$K \quad f_{mk}$$

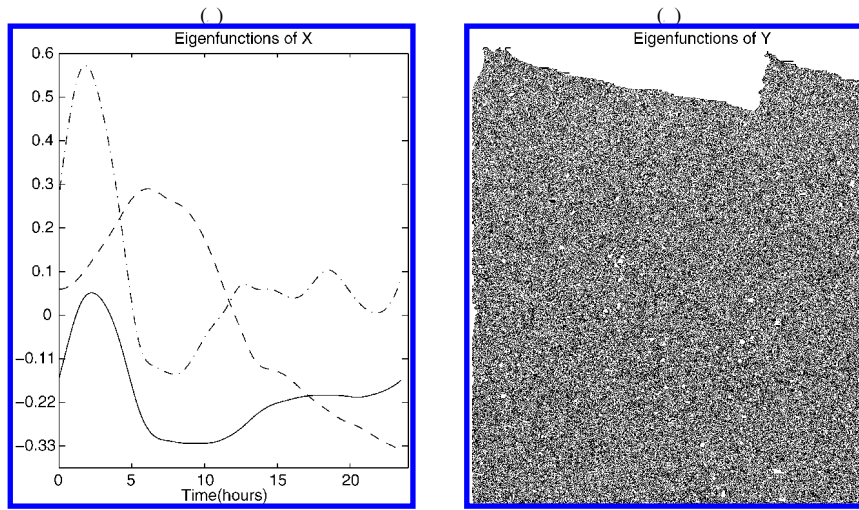




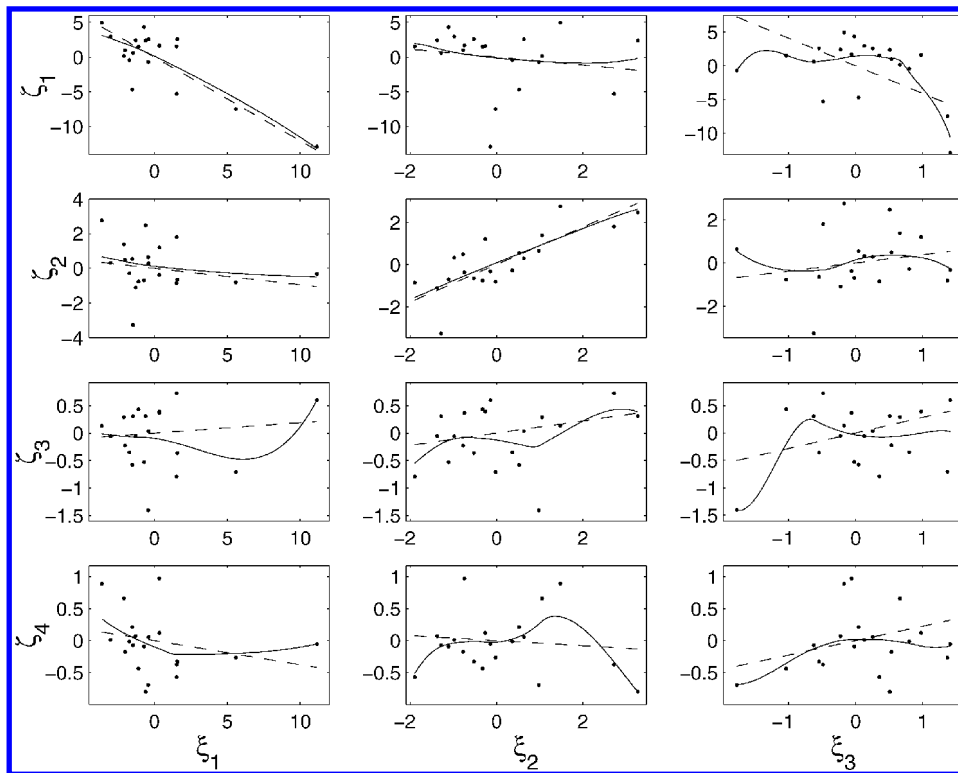


### 7. APPLICATION TO GENE EXPRESSION TIME COURSE DATA





$\xi_1 = \frac{1}{\sqrt{2}}(x_1 - x_2)$      $\xi_2 = \frac{1}{\sqrt{2}}(x_1 + x_2)$      $\xi_3 = \frac{1}{\sqrt{2}}(x_1 - x_2)$      $\xi_4 = \frac{1}{\sqrt{2}}(x_1 + x_2)$      $K = \dots$   
 $\xi_1 = \frac{1}{\sqrt{2}}(x_1 - x_2)$      $\xi_2 = \frac{1}{\sqrt{2}}(x_1 + x_2)$      $\xi_3 = \frac{1}{\sqrt{2}}(x_1 - x_2)$      $\xi_4 = \frac{1}{\sqrt{2}}(x_1 + x_2)$      $M = \dots$



$\xi_1 = \frac{1}{\sqrt{2}}(x_1 - x_2)$      $\xi_2 = \frac{1}{\sqrt{2}}(x_1 + x_2)$      $\xi_3 = \frac{1}{\sqrt{2}}(x_1 - x_2)$      $\xi_4 = \frac{1}{\sqrt{2}}(x_1 + x_2)$      $K = \dots$   
 $\xi_1 = \frac{1}{\sqrt{2}}(x_1 - x_2)$      $\xi_2 = \frac{1}{\sqrt{2}}(x_1 + x_2)$      $\xi_3 = \frac{1}{\sqrt{2}}(x_1 - x_2)$      $\xi_4 = \frac{1}{\sqrt{2}}(x_1 + x_2)$      $M = \dots$





$(j, l) = 1, \dots, n_i$  for  $i = 1, \dots, M$ . Let  $U_i = (U_{i1}, \dots, U_{in_i})^T$ ,  $X_i = (\mu_X(s_{i1}), \dots, \mu_X(s_{in_i}))^T$ ,  $U_{ik} = (U_{ik}(s_{i1}), \dots, U_{ik}(s_{in_i}))^T$ ,  $U_{ik} = \sum_{j=1}^{n_i} U_{ij} U_{jl} = G_X(s_{ij}, s_{il}) + X_j^T X_l$ ,  $j, l = 1, \dots, n_i$ ,  $j \neq l$ . Let  $\hat{U}_{ik} = \hat{U}_{ik}^T \hat{\Sigma}_{U_i}^{-1} (U_i - \hat{X}_i)$ ,  $(j, l) = 1, \dots, n_i$ ,  $\hat{\Sigma}_{U_i} = (\hat{\Sigma}_{U_i})_{j,l} = \hat{G}_X(s_{ij}, s_{il}) + \hat{Y}(s_{ij})$ ,  $j, l = 1, \dots, n_i$ ,  $j \neq l$ . Let  $\hat{I}_{ik}^I = \sum_{j=1}^{n_i} (U_{ij} - \hat{\mu}_X(t_{ij})) \hat{k}(s_{ij})(s_{ij} - s_{i,j-})$ ,  $\hat{I}_{im}^I = \sum_{j=1}^{n_i} (V_{ij} - \hat{\mu}_Y(t_{ij})) \hat{k}(t_{ij})(t_{ij} - t_{i,j-})$ .

APPENDIX: ESTIMATION PROCEDURES AND ASSUMPTIONS

A.1 Eigenrepresentations and Estimating Functional Principal Component Scores

$k, k = 1, \dots, \infty$ . Let  $(A_{G_X} f)(s) = \int_{\mathcal{H}} G_X(s, t) f(t) dt$ ,  $f \in L(\mathcal{S})$ ,  $(A_{G_X} f)(t) = \int_{\mathcal{H}} f(s) G_X(s, t) ds$ ,  $\int_{\mathcal{H}} \int_{\mathcal{H}} f_j(s) k(s) ds = \int_{\mathcal{H}} f_j(s) k(s) ds$ ,  $j, k = 1, \dots, \infty$ ,  $j \neq k$ .

$\{\mu_X, j\}$ ,  $\{\mu_Y, k\}$ ,  $G_X(s, s) = \sum_{k=1}^{\infty} k \times k(s) k(s)$ ,  $G_Y(t, t) = \sum_{m=1}^{\infty} m \times m(t) m(t)$ .

$$X(s) = \mu_X(s) + \sum_{j=1}^{\infty} j j(s) \quad (1)$$

$$Y(t) = \mu_Y(t) + \sum_{k=1}^{\infty} k k(t).$$

$\{k, k = 1, \dots, \infty\}$ ,  $\{m, m = 1, \dots, \infty\}$ . Let  $\hat{X}(s) = \hat{\mu}_X(s) + \sum_{j=1}^{\infty} \hat{I}_{ij}^I j(s)$ ,  $\hat{Y}(t) = \hat{\mu}_Y(t) + \sum_{k=1}^{\infty} \hat{I}_{im}^I k(t)$ .

$$U_{ij} = X_i(s_{ij}) + ij = \mu_X(s_{ij}) + \sum_{k=1}^{\infty} ik k(s_{ij}) + ij, \quad s_{ij} \in \mathcal{S}, \quad 1 \leq i \leq n, \quad 1 \leq j \leq n_i. \quad (2)$$

$$V_{il} = Y_i(t_{il}) + il = \mu_Y(t_{il}) + \sum_{m=1}^{\infty} im m(t_{il}) + il, \quad s_{il} \in \mathcal{T}, \quad 1 \leq i \leq n, \quad 1 \leq l \leq m_i. \quad (3)$$

$M$ ,  $U_i = (U_{i1}, \dots, U_{in_i})^T$ ,  $X_i = (\mu_X(s_{i1}), \dots, \mu_X(s_{in_i}))^T$ ,  $U_{ik} = (U_{ik}(s_{i1}), \dots, U_{ik}(s_{in_i}))^T$ ,  $U_{ik} = \sum_{j=1}^{n_i} U_{ij} U_{jl} = G_X(s_{ij}, s_{il}) + X_j^T X_l$ ,  $j, l = 1, \dots, n_i$ ,  $j \neq l$ . Let  $\hat{U}_{ik} = \hat{U}_{ik}^T \hat{\Sigma}_{U_i}^{-1} (U_i - \hat{X}_i)$ ,  $(j, l) = 1, \dots, n_i$ ,  $\hat{\Sigma}_{U_i} = (\hat{\Sigma}_{U_i})_{j,l} = \hat{G}_X(s_{ij}, s_{il}) + \hat{Y}(s_{ij})$ ,  $j, l = 1, \dots, n_i$ ,  $j \neq l$ . Let  $\hat{I}_{ik}^I = \sum_{j=1}^{n_i} (U_{ij} - \hat{\mu}_X(t_{ij})) \hat{k}(s_{ij})(s_{ij} - s_{i,j-})$ ,  $\hat{I}_{im}^I = \sum_{j=1}^{n_i} (V_{ij} - \hat{\mu}_Y(t_{ij})) \hat{k}(t_{ij})(t_{ij} - t_{i,j-})$ .

$\hat{I}_{ik}^I = \sum_{j=1}^{n_i} (U_{ij} - \hat{\mu}_X(t_{ij})) \hat{k}(s_{ij})(s_{ij} - s_{i,j-})$ ,  $\hat{I}_{im}^I = \sum_{j=1}^{n_i} (V_{ij} - \hat{\mu}_Y(t_{ij})) \hat{k}(t_{ij})(t_{ij} - t_{i,j-})$ .

$$\hat{I}_{ik}^I = \sum_{j=1}^{n_i} (U_{ij} - \hat{\mu}_X(t_{ij})) \hat{k}(s_{ij})(s_{ij} - s_{i,j-}), \quad (4)$$

$$\hat{I}_{im}^I = \sum_{j=1}^{n_i} (V_{ij} - \hat{\mu}_Y(t_{ij})) \hat{k}(t_{ij})(t_{ij} - t_{i,j-}), \quad (5)$$

$\hat{I}_{ik}^I = \sum_{j=1}^{n_i} (U_{ij} - \hat{\mu}_X(t_{ij})) \hat{k}(s_{ij})(s_{ij} - s_{i,j-})$ ,  $\hat{I}_{im}^I = \sum_{j=1}^{n_i} (V_{ij} - \hat{\mu}_Y(t_{ij})) \hat{k}(t_{ij})(t_{ij} - t_{i,j-})$ .

A.2 Estimation Procedures

$$\sum_{i=1}^n \sum_{j=1}^{n_i} K\left(\frac{s_{ij} - s}{b_X}\right) \{U_{ij} - X_i(s_{ij}) - X(s - s_{ij})\} \quad (6)$$

$$\hat{\mu}_X(s) = \hat{X}(s) = \sum_{i=1}^n \sum_{j=1}^{n_i} K\left(\frac{s_{ij} - s}{b_X}\right) \{U_{ij} - \hat{\mu}_X(s_{ij})\} \quad (7)$$

$$\sum_{i=1}^n \sum_{j=1}^{n_i} K\left(\frac{s_{ij} - s}{h_X}, \frac{s_{ij} - s}{h_X}\right) \times \{X_i(s_{ij}, s_{ij}) - f(X(s, s), (s_{ij}, s_{ij}))\}, \quad (8)$$

$$f(X(s, s), (s_{ij}, s_{ij})) = X + X(s - s_{ij}) + X(s - s_{ij}), \quad \hat{X} = (\hat{X}, \hat{X}, \hat{X})^T, \quad \hat{G}_X(s, s) = \hat{X}(s, s), \quad \hat{K} = \hat{K}, \quad \hat{h}_X = \hat{h}_X, \quad \hat{j} = \hat{j}, \quad \hat{X} = \hat{X}, \quad \hat{G}_X = \hat{G}_X, \quad \hat{X}(s) = \hat{X}(s), \quad \{G_X(s, s) + X\} = \hat{V}_X(s)$$

$$h_{X^*}^* = \{a = \int_{S^*} \{s \mid s \in S\}, b = \int_{S^*} \{s \mid s \in S\} \mid S\} = b - a, \quad S = [a + |S|/2, b - |S|/2].$$

$$\hat{X} = \int \{\hat{V}_X(s) - \tilde{G}_X(s)\} ds / |S| \quad (1)$$

$$\hat{X} > \dots \hat{X} = \int \{k, k\}_{k \geq \dots} \mathbb{E} \{ \dots \} \{k, k\}$$

$$\begin{aligned}
 & \{ \dots \} \\
 & \sum_{k=1}^K [f_k''(k)h_k + n^{-1} \{ \dots (Y|k) \} / p_k^{-1}(k) \times \\
 & \quad h_k^{-1}] \rightarrow \\
 & \sum_{k=1}^K \sum_{m=1}^M [f_{mk}''(k) m(t)h_{mk} + n^{-1} \{ \dots (m|k) \} / \times \\
 & \quad p_k^{-1}(k) | m(t)h_{mk}^{-1}] \rightarrow \dots
 \end{aligned}$$

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# Supplement for “Functional Additive Models”

## 1. NOTATIONS AND AUXILIARY RESULTS

Covariance operators are denoted by  $\mathcal{G}_X$ ,  $\widehat{\mathcal{G}}_X$ , generated by kernels  $G_X$ ,  $\widehat{G}_X$ ; i.e.,  $\mathcal{G}_X(f) = \int_{\mathcal{T}} G_X(s, t)f(s)ds$ ,  $\widehat{\mathcal{G}}_X(f) = \int_{\mathcal{T}} \widehat{G}_X(s, t)f(s)ds$  for any  $f \in L^2(\mathcal{T})$ . Define

$$\begin{aligned} D_X &= \int_{\mathcal{T}^2} \{\widehat{G}_X(s, t) - G_X(s, t)\}^2 dsdt, & \delta_k^X &= \min_{1 \leq j \leq k} (\lambda_j - \lambda_{j+1}), \\ K_0 &= \inf\{j \geq 1 : \lambda_j - \lambda_{j+1} \leq 2D_X\} - 1, & \pi_k^X &= 1/\lambda_k + 1/\delta_k^X. \end{aligned} \quad (32)$$

Let  $K = K(n)$  denote the numbers of leading eigenfunctions included to approximate  $X$  as sample size  $n$  varies; i.e.,  $\widehat{X}_i(s) = \widehat{\mu}_X(s) + \sum_{k=1}^K \widehat{\xi}_{ik} \widehat{\phi}_k(s)$ . Analogously, define the quantities  $\mathcal{G}_Y$ ,  $\widehat{\mathcal{G}}_Y$ ,  $D_Y$ ,  $\delta_m^Y$ ,  $\pi_m^Y$ ,  $M_0$  and  $M$  for the process  $Y$ , for the case of functional responses. The following lemma gives the weak uniform convergence rates for the estimators of the FPCs, setting the stage for the subsequent developments. The proof is in Section 2.

**Lemma 1** Under  $A$   $A$   $C$   $C$   $\rightarrow nd$   $C$

$$\sup_{t \in \mathcal{S}} |\widehat{\mu}_X(s) - \mu_X(s)| = O_p\left(\frac{1}{\sqrt{nb_X}}\right), \quad \sup_{s_1, s_2 \in \mathcal{S}} |\widehat{G}_X(s_1, s_2) - G_X(s_1, s_2)| = O_p\left(\frac{1}{\sqrt{nh_X^2}}\right), \quad (33)$$

$\rightarrow nd$   $\rightarrow s$   $\rightarrow$  consequence  $\widehat{\sigma}_X^2 - \sigma_X^2 = O_p(n^{-1/2}h_X^{-2} + n^{-1/2}h_X^{*-1})$  Considering eigenvectors  $\lambda_k$  of,  $u$  tip icity one  $\widehat{\phi}_k$   $c_n$  e chosen such that

$$P\left(\sup_{1 \leq k \leq K_0} |\widehat{\lambda}_k - \lambda_k| \leq D_X\right) = 1, \quad \sup_{s \in \mathcal{S}} |\widehat{\phi}_k(s) - \phi_k(s)| = O_p\left(\frac{\pi_k^X}{\sqrt{nh_X^2}}\right), \quad k = 1, \dots, K_0, \quad (34)$$

here  $D_X$   $\pi_k^X$   $\rightarrow nd$   $K_0$   $\rightarrow$  re de ned in

Analogous under  $B$   $B$   $C$   $C$   $\rightarrow nd$   $C$

$$\sup_{t \in \mathcal{T}} |\widehat{\mu}_Y(t) - \mu_Y(t)| = O_p\left(\frac{1}{\sqrt{nb_Y}}\right), \quad \sup_{t_1, t_2 \in \mathcal{T}} |\widehat{G}_Y(t_1, t_2) - G_Y(t_1, t_2)| = O_p\left(\frac{1}{\sqrt{nh_Y^2}}\right), \quad (35)$$

$\rightarrow nd$   $\rightarrow s$   $\rightarrow$  consequence  $\widehat{\sigma}_Y^2 - \sigma_Y^2 = O_p(n^{-1/2}h_Y^{-2} + n^{-1/2}h_Y^{*-1})$  Considering eigenvectors  $\rho_m$  of,  $u$  tip icity one  $\widehat{\psi}_m$   $c_n$  e chosen such that

$$P\left(\sup_{1 \leq m \leq M_0} |\widehat{\rho}_m - \rho_m| \leq D_Y\right) = 1, \quad \sup_{t \in \mathcal{T}} |\widehat{\psi}_m(t) - \psi_m(t)| = O_p\left(\frac{\pi_m^Y}{\sqrt{nh_Y^2}}\right), \quad m = 1, \dots, M_0, \quad (36)$$

here  $D_{\mathbf{Y}} \pi_{\mathbf{k}}^{\mathbf{y}}$  and  $M_0$  are defined analogously to for process  $Y$

Recall that  $\|f\|_{\infty} = \sup_{\mathbf{x} \in \mathcal{A}} |f(\mathbf{x})|$  for an arbitrary function  $f$  with support  $\mathcal{A}$ , and  $\|g\| = \sqrt{\int_{\mathcal{A}} g^2(t) dt}$  for any  $g \in L^2(\mathcal{A})$  and define

$$\begin{aligned} \theta_{\mathbf{ik}}^{(1)} &= c_1 \|X_{\mathbf{i}}\| + c_2 \|X_{\mathbf{i}} X'_{\mathbf{i}}\|_{\infty} + c_3, & Z_{\mathbf{k}}^{(1)} &= \sup_{s \in \mathcal{S}} |\hat{\phi}_{\mathbf{k}}(s) - \phi_{\mathbf{k}}(s)|, \\ \theta_{\mathbf{ik}}^{(2)} &= \mathbf{1} + \|\phi_{\mathbf{k}} \phi'_{\mathbf{k}}\|_{\infty}, & Z_{\mathbf{k}}^{(2)} &= \sup_{s \in \mathcal{S}} |\hat{\mu}_{\mathbf{x}}(s) - \mu_{\mathbf{x}}(s)|, \\ \theta_{\mathbf{ik}}^{(3)} &= c_4 \|X_{\mathbf{i}}\|_{\infty} + c_5 \|X'_{\mathbf{i}}\|_{\infty} + c_6, & Z_{\mathbf{k}}^{(3)} &= \|\phi'_{\mathbf{k}}\|_{\infty}, \\ \theta_{\mathbf{ik}}^{(4)} &= \left| \sum_{j=2}^{n_i} \epsilon_{ij} \phi_{\mathbf{k}}(s_{ij})(s_{ij} - s_{i,j-1}) \right|, & Z_{\mathbf{k}}^{(4)} &\equiv \mathbf{1}, \\ \theta_{\mathbf{ik}}^{(5)} &= \sum_{j=2}^{n_i} |\epsilon_{ij}| (s_{ij} - s_{i,j-1}), & & k \end{aligned}$$

**Lemma 2** For  $\theta_{\mathbf{ik}}^{(\cdot)}$ ,  $Z_{\mathbf{k}}^{(\cdot)}$ ,  $\vartheta_{\mathbf{im}}^{(\cdot)}$  and  $Q_{\mathbf{m}}^{(\cdot)}$  as defined in (38)

$$|\hat{\xi}_{\mathbf{ik}}^I - \xi_{\mathbf{ik}}| \leq \sum_{=1}^5 \theta_{\mathbf{ik}}^{(\cdot)} Z_{\mathbf{k}}^{(\cdot)}, \quad |\hat{\zeta}_{\mathbf{im}}^I - \zeta_{\mathbf{im}}| \leq \sum_{=1}^5 \vartheta_{\mathbf{im}}^{(\cdot)} Q_{\mathbf{m}}^{(\cdot)}. \quad (39)$$

The proof is in Section 2. In the sequel we suppress the superscript  $I$  in the FPC estimates  $\hat{\xi}_{\mathbf{ik}}^I$  and  $\hat{\zeta}_{\mathbf{im}}^I$ .

Recall that the sequences of bandwidths  $h_{\mathbf{k}}$  and  $h_{\mathbf{mk}}$  are employed to obtain the estimates  $\hat{f}_{\mathbf{k}}$  and  $\hat{f}_{\mathbf{mk}}$  for the regression functions  $f_{\mathbf{k}}$  and  $f_{\mathbf{mk}}$ , and that the density of  $\xi_{\mathbf{k}}$  is denoted by  $p_{\mathbf{k}}$ . Define

$$\begin{aligned} \theta_{\mathbf{k}}(x) &= p_{\mathbf{k}}(x) \left\{ \frac{\pi_{\mathbf{k}}^{\mathbf{x}}}{\sqrt{nh_{\mathbf{k}}^2}} + \frac{1}{\sqrt{nb_{\mathbf{x}}}} + \sqrt{\frac{1}{b_{\mathbf{x}}^*}} \right\}, \\ \vartheta_{\mathbf{mk}}(x) &= p_{\mathbf{k}}(x) \left\{ \frac{\pi_{\mathbf{m}}^{\mathbf{y}}}{\sqrt{nh_{\mathbf{y}}^2}} + \frac{1}{\sqrt{nb_{\mathbf{y}}}} + \sqrt{\frac{1}{b_{\mathbf{y}}^*}} \right\}. \end{aligned} \quad (40)$$

The weak convergence rates  $\tilde{\theta}_{\mathbf{k}}$  and  $\tilde{\vartheta}_{\mathbf{mk}}$  of the regression function estimators  $\hat{f}_{\mathbf{k}}(x)$  and  $\hat{f}_{\mathbf{mk}}(x)$  (see Theorem 1) are as follows,

$$\begin{aligned} \tilde{\theta}_{\mathbf{k}}(x) &= \frac{\theta_{\mathbf{k}}(x)}{h_{\mathbf{k}}} + \frac{1}{2} |f_{\mathbf{k}}''(x)| h_{\mathbf{k}}^2 + \sqrt{\frac{\text{var}(Y|x) \|K_1\|^2}{p_{\mathbf{k}}(x) n h_{\mathbf{k}}}}, \\ \tilde{\vartheta}_{\mathbf{mk}}(x) &= \frac{\theta_{\mathbf{k}}(x)}{h_{\mathbf{mk}}} + \vartheta_{\mathbf{mk}}(x) + \frac{1}{2} |f_{\mathbf{mk}}''(x)| h_{\mathbf{mk}}^2 + \sqrt{\frac{\text{var}(\zeta_{\mathbf{m}}|x) \|K_1\|^2}{p_{\mathbf{k}}(x) n h_{\mathbf{mk}}}}. \end{aligned} \quad (41)$$

Considering the predictions  $\hat{E}(Y|X)$  for the scalar response case and  $\hat{E}\{Y(t)|X\}$  for the functional response case, the numbers of eigenfunctions  $K$  and  $M$  used for approximating the infinite dimensional processes  $X$  and  $Y$  generally tend to infinity as the sample size  $n$  increases. We require  $K \leq K_0$  and  $M \leq M_0$  in (A6). Since it follows from (35) that  $K_0 \rightarrow \infty$ , as long as all eigenvalues  $\lambda_j$  are of multiplicity 1, and analogously for  $M_0$ , this is not a strong restriction. Denote the set of positive integers by  $\mathcal{N}$  and  $\mathcal{N}_{\mathbf{k}} = \{1, \dots, k\}$ . Convergence rates  $\theta_n^*$  and  $\vartheta_n^*$  for the predictions (20) and

(21) are as follows,

$$\begin{aligned} \theta_n^* &= \sum_{k=1}^K \left\{ \frac{\theta_k(\xi_k)}{h_k} + \frac{1}{2} |f_k''(\xi_k)| h_k^2 + \sqrt{\frac{\text{var}(Y|\xi_k) \|K_1\|^2}{p_k(\xi_k) n h_k}} \right\} + \left| \sum_{k \geq K+1} f_k(\xi_k) \right|, \quad (42) \\ \vartheta_n^* &= \sum_{k=1}^K \sum_{m=1}^M \left\{ \left( \frac{\theta_k(\xi_k)}{h_{mk}} + \vartheta_{mk}(\xi_k) \right) |\psi_m(t)| + \frac{1}{2} |f_{mk}''(\xi_k)| |\psi_m(t)| h_{mk}^2 + \sqrt{\frac{\text{var}(\zeta_m|\xi_k) \|K_1\|^2}{p_k(\xi_k) n h_k}} |\psi_m(t)| \right. \\ &\quad \left. + \frac{\pi_m^y |f_{mk}(\xi_k)|}{h_{mk}} \right\} \end{aligned}$$

Noting  $\hat{\xi}_{ik} = \hat{\eta}_{ik} + \hat{\tau}_{ik}$ , one finds

$$|\hat{\xi}_{ik} - \xi_{ik}| \leq \{|\hat{\eta}_{ik} - \tilde{\eta}_{ik}| + |\tilde{\eta}_{ik} - \xi_{ik}| + |\hat{\tau}_{ik}|\}. \quad (43)$$

Without loss of generality, assume  $\|\phi_k\|_\infty \geq 1$ ,  $\|\phi'_k\|_\infty \geq 1$ ,  $\|X_i\|_\infty \geq 1$  and  $\|X'_i\|_\infty \geq 1$ .

For  $\theta_{ik}^{(\ell)}$  and  $Z_k^{(\ell)}$  (37),  $\ell = 1, \dots, 5$ , the first term on the r.h.s. of (43) is bounded by

$$\begin{aligned} & \left\{ \sum_{j=2}^{n_i} [|X_i(s_{ij}) - \hat{\mu}(s_{ij})| \cdot |\hat{\phi}_k(s_{ij}) - \phi_k(s_{ij})| + |\hat{\mu}(s_{ij}) - \mu(s_{ij})| \cdot |\phi_k(s_{ij})|] (s_{ij} - s_{i,j-1}) \right\} \\ & \leq \left\{ \sum_{j=1}^{n_i} [|X_i(s_{ij})| + |\mu(s_{ij})| + 1]^2 (s_{ij} - s_{i,j-1}) \right\}^{1/2} \left\{ \sum_{j=2}^{n_i} [\hat{\phi}_k(s_{ij}) - \phi_k(s_{ij})]^2 (s_{ij} - s_{i,j-1}) \right\}^{1/2} \\ & \quad + \left\{ \sum_{j=1}^{n_i} [\hat{\mu}(s_{ij}) - \mu(s_{ij})]^2 (s_{ij} - s_{i,j-1}) \right\}^{1/2} \left\{ \sum_{j=2}^{n_i} \phi_k^2(s_{ij}) (s_{ij} - s_{i,j-1}) \right\}^{1/2} \\ & \leq \theta_{ik}^{(1)} Z_k^{(1)} + \theta_{ik}^{(2)} Z_k^{(2)}. \end{aligned}$$

The second term on the r.h.s. of (43) has the upper bound

$$|\tilde{\eta}_{ij} - \xi_{ik}| \leq \|(X_i + \mu)' \phi_k + (X_i + \mu) \phi'_k\|_\infty \cdot \mathbf{x} \leq \theta_{ik}^{(3)} Z_k^{(3)}.$$

From the above, the third term on the r.h.s. of (43) is bounded by  $(\theta_{ik}^{(4)} Z_k^{(4)} + \theta_{ik}^{(5)} Z_k^{(5)})$ .

□

*Proof of Theore* For simplicity, denote " $\sum_{i=1}^n$ " by " $\sum_i$ ",  $w_i = K_1 \{(x - \xi_{ik})/h_k\}/(nh_k)$ ,  $\hat{w}_i = K_1 \{(x - \hat{\xi}_{ik})/h_k\}/(nh_k)$ , and write  $\theta_k = \theta_k(x)$ . From (12), the local linear estimator  $\hat{f}_k(x)$  of the regression function  $f_k(x)$  can be explicitly written as

$$\hat{f}_k(x) = \frac{\sum_i \hat{w}_i Y_i}{\sum_i \hat{w}_i} - \frac{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x)}{\sum_i \hat{w}_i} \hat{f}'_k(x), \quad (44)$$

where

$$\hat{f}'_k(x) = \frac{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x) Y_i - \{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x) \sum_i \hat{w}_i Y_i\} / \sum_i \hat{w}_i}{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x)^2 - \{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x)\}^2 / \sum_i \hat{w}_i}. \quad (45)$$

Let  $\tilde{f}_k(x)$  be a hypothetical estimator, obtained by substituting the true values  $w_i$  and  $\xi_{ik}$  for  $\hat{w}_i$ ,  $\hat{\xi}_{ik}$  in (44) and (45). To evaluate  $|\hat{f}_k(x) - \tilde{f}_k(x)|$ , one has to quantify the



orders of the differences

$$\begin{aligned} D_1 &= \sum_{\mathbf{i}} (\hat{w}_{\mathbf{i}} - w_{\mathbf{i}}), & D_2 &= \sum_{\mathbf{i}} (\hat{w}_{\mathbf{i}} - w_{\mathbf{i}}) Y_{\mathbf{i}}, \\ D_3 &= \sum_{\mathbf{i}} (\hat{w}_{\mathbf{i}} \hat{\xi}_{\mathbf{i}\mathbf{k}} - w_{\mathbf{i}} \xi_{\mathbf{i}\mathbf{k}}), & D_4 &= \sum_{\mathbf{i}} (\hat{w}_{\mathbf{i}} \hat{\xi}_{\mathbf{i}\mathbf{k}}^2 - w_{\mathbf{i}} \xi_{\mathbf{i}\mathbf{k}}^2). \end{aligned}$$

Considering  $D_1$ , without loss of generality, assume the compact support of  $K_1$  is  $[-1, 1]$ .

Since  $K_1$  is Lipschitz continuous on its support,

$$D_1 \leq \frac{c}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} |\hat{\xi}_{\mathbf{i}\mathbf{k}} - \xi_{\mathbf{i}\mathbf{k}}| \{I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) + I(|x - \hat{\xi}_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}})\}, \quad (46)$$

for some  $c > 0$ , where  $I(\cdot)$  is an indicator function. Lemma 2 implies for the first term on the r.h.s. of (46)

$$\frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} |\hat{\xi}_{\mathbf{i}\mathbf{k}} - \xi_{\mathbf{i}\mathbf{k}}| I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) \leq \sum_{\ell=1}^5 Z_{\mathbf{k}}^{(\ell)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(\ell)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}).$$

Applying the central limit theorem for a random number of summands (Billingsley, 1995, page 380), observing  $\sum_{\mathbf{i}} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) / (nh_{\mathbf{k}}) \xrightarrow{p} 2p_{\mathbf{k}}(x)$ , one finds

$$\frac{1}{nh_{\mathbf{k}}} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(\ell)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) \xrightarrow{p} 2p_{\mathbf{k}}(x) E(\theta_{\mathbf{i}\mathbf{k}}^{(\ell)}), \quad (47)$$

provided that  $E(\theta_{\mathbf{i}\mathbf{k}}^{(\ell)}) < \infty$  for  $\ell = 1, \dots, 5$ . Note that  $E\theta_{\mathbf{i}\mathbf{k}}^{(1)} < \infty$ ,  $E\theta_{\mathbf{i}\mathbf{k}}^{(3)} < \infty$  by (A4),

$E\theta_{\mathbf{i}\mathbf{k}}^{(4)} \leq 2\sigma_{\mathbf{X}} \sqrt{\mathbf{X}^*}$  and  $E\theta_{\mathbf{i}\mathbf{k}}^{(5)} \leq |S|\sigma_{\mathbf{X}}$  by the Cauchy-Schwarz inequality. Then

$$\begin{aligned} Z_{\mathbf{k}}^{(1)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(1)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{\pi_{\mathbf{k}}^{\mathbf{x}}}{\sqrt{nh_{\mathbf{X}}^2} h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}, \\ Z_{\mathbf{k}}^{(2)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(2)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{1}{\sqrt{nb_{\mathbf{X}}} h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}, \\ Z_{\mathbf{k}}^{(3)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(3)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{\|\phi_{\mathbf{k}}\|_{\infty} \mathbf{X}^*}{h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}, \\ Z_{\mathbf{k}}^{(4)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(4)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{\sqrt{\mathbf{X}^*}}{h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}, \\ Z_{\mathbf{k}}^{(5)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(5)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{\pi_{\mathbf{k}}^{\mathbf{x}}}{\sqrt{nh_{\mathbf{X}}^2} h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}. \end{aligned} \quad (48)$$

We now obtain  $(nh_k^2)^{-1} \sum_i |\hat{\xi}_{ik} - \xi_{ik}| I(|x - \xi_{ik}| \leq h_k) = O_p(\theta_k h_k^{-1})$ . The asymptotic rate of the second term can be derived analogously, observing

$$\frac{1}{nh_k} \sum_i I(|x - \hat{\xi}_{ik}| \leq h_k) \leq \frac{1}{nh_k} \sum_i \{I(|x - \xi_{ik}| \leq 2h_k) + I(\sum_{l=1}^5 \theta_{ik}^{(l)} Z_k^{(l)} > h_k)\} \xrightarrow{p} 4p_k(x),$$

leading to  $(nh_k^2)^{-1} \sum_i |\hat{\xi}_{ik} - \xi_{ik}| I(|x - \hat{\xi}_{ik}| \leq h_k) = O_p(\theta_k h_k^{-1})$ . Then  $D_1 = O_p(\theta_k h_k^{-1})$  follows.

Analogously, one shows  $D_2 = O_p(\theta_k h_k^{-1})$ , applying the Cauchy-Schwarz inequality for  $\theta_{ik}^{(\ell)}$ ,  $\ell = 1, 3$ , and observing the independence between  $Y_i$  and  $\theta_{ik}^{(\ell)}$  for  $\ell = 2, 4, 5$ , given the moment condition (A4). For  $D_3$ , observe

$$D_3 = \sum_i \{(\hat{w}_i - w_i)\xi_{ik} + (\hat{w}_i - w_i)(\hat{\xi}_{ik} - \xi_{ik}) + w_i(\hat{\xi}_{ik} - \xi_{ik})\} \equiv D_{31} + D_{32} + D_{33}.$$

Then  $D_{31} = O_p(\theta_k h_k^{-1})$ , analogously to  $D_1$ . It is easy to see that  $D_{32} = o_p(D_{31})$ . Since  $D_{33} \leq c \sum_{l=1}^5 Z_k^{(l)} (nh_k)^{-1} \sum_i \theta_{ik}^{(l)} I(|x - \xi_{ik}| \leq h_k)$  for some  $c > 0$ , one also has  $D_{33} = o_p(D_{31})$ . This results in  $D_3 = O_p(\theta_k h_k^{-1})$ . Observing  $|\hat{\xi}_{ik}^2 - \xi_{ik}^2| \leq |\hat{\xi}_{ik} - \xi_{ik}| \cdot |\xi_{ik}| + (\hat{\xi}_{ik} - \xi_{ik})^2$ , one can show  $D_4 = O_p(\theta_k h_k^{-1})$ , using similar arguments as for  $D_3$ , and  $E\xi_{ik}^4 < \infty$  from (A4). Combining the results for  $D$ ,  $\ell = 1, \dots, 4$ , and applying Slutsky's Theorem leads to  $|\hat{f}_k(x) - \tilde{f}_k(x)| = O_p(\theta_k h_k^{-1})$ . Using (A5), and applying standard asymptotic results for the hypothetical local linear smoother  $\tilde{f}_k(x)$  completes the proof of (18).

To derive (19), additionally one only needs to consider  $\sum_i (\hat{w}_i \hat{\zeta}_{im} - w_i \zeta_{im}) = \sum_i \{(\hat{w}_i - w_i)\zeta_{im} + (\hat{w}_i - w_i)(\hat{\zeta}_{im} - \zeta_{im}) + w_i(\hat{\zeta}_{im} - \zeta_{im})\}$ , where the third term yields an extra term of order  $O_p(\vartheta_{mk})$  by observing

$$|\sum_i w_i(\hat{\zeta}_{im} - \zeta_{im})| \leq \sum_{m=1}^5 Q_m^{(\cdot)} \sum_i w_i \vartheta_{im}^{(\cdot)} \leq \frac{1}{nh_{mk}} \sum_{m=1}^5 Q_m^{(\cdot)} \sum_i \vartheta_{im}^{(\cdot)} I(|x - \xi_{ik}| \leq h_{mk})$$

Similar arguments as above complete this derivation.  $\square$

*Proof of Theore*

Using (A7), the derivation of  $\theta_n^*$  in (42) is straightforward,

following the above arguments. To obtain (21), note that

$$\begin{aligned}
& \widehat{E}\{Y(t)|X\} - E\{Y(t)|X\} \\
\leq & \sum_{k=1}^K \sum_{m=1}^M |\widehat{f}_{mk}(\xi_k)\widehat{\psi}_m(t) - f_{mk}(\xi_k)\psi_m(t)| + \left| \sum_{k \geq K+1} \sum_{m \geq M+1} f_{mk}(\xi_k)\psi_m(t) \right| \\
\leq & \sum_{k=1}^K \sum_{m=1}^M [|\widehat{f}_{mk}(\xi_k) - f_{mk}(\xi_k)|\{|\psi_m(t)| + |\widehat{\psi}_m(t) - \psi_m(t)|\} + |f_{mk}(\xi_k)| \cdot |\widehat{\psi}_m(t) - \psi_m(t)|] \\
& + \left| \sum_{(\mathbf{k}, \mathbf{m}) \in \mathcal{N}^2 \setminus \mathcal{N}_K \times \mathcal{N}_M} f_{mk}(\xi_k)\psi_m(t) \right|.
\end{aligned}$$

This implies the convergence rate  $\vartheta_n^*$  in (42).